Market Responsiveness to Corporate Social Responsibility

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Abstract

Market reaction to surprises in earnings announcements has long been used to measure the quality of the information content of the announcement, and studies have explored various factors affecting the response. This study adds to this body of research by factoring in the level of corporate social responsibility (CSR) exhibited by the firm and employs a relatively new measure of a company's level of CSR, rankings published by JUST Capital. This study hypothesizes that financial information reported by higher ranked companies is weighed more heavily by investors than those reported by non-ranked or lower-ranked companies. Using earnings response coefficients as a measure of the perceived quality of the financial information reported by the firms on the marked response. The results provide direct support of the hypotheses, indicating that the market reacts more strongly to earnings surprises for firms with high JUST rankings than for unranked firms as well as reacting stronger for higher ranked firms relative to lower ranked firms. We conclude from this study that investors do take into account the level of a corporation's social responsibility when evaluating earnings announcements. This result contributes new insights into the impact of a firm's CSR in terms of the perceived quality of a firm's financial reporting.

Keywords: corporate social responsibility, market reaction, earnings quality, earnings response coefficient, JUST capital ranking

1. Introduction

Previous literature has identified a positive relation between a corporation's social responsibility (CSR) performance and its financial performance (Orlitzky et al., 2003: Friede et al., 2015). However, many of these studies suffer from issues related to the definition and measurability of attributes such as social and environmental responsibility (Pava and Krausz, 1996) as well as from a variety of self-selection and identification issues (Aupperle et al., 1985; McWilliams and Siegel, 2000).

Recently, a non-for profit organization known as JUST Capital began issuing an annual ranking of firms in the Russel 1000 based on environmental, social and governance (ESG) metrics.¹ These rankings have gained traction in the investment community. An index tracking the top 50% of the Russel 1000 weighted by JUST rankings (known as the JUST 500 or JULCD) began trading in November 2016, and in June 2018, Goldman Sachs Asset Management launched an Exchange Traded Fund (ETF) based on this index.

In this paper, we examine whether there is a relation between the rankings published by JUST Capital and the perceived quality of financial reporting for these firms. Using earnings response coefficients (ERCs) to measure the perceived quality of information contained in the earnings report, this study tests to see if firms which are ranked by JUST have larger ERCs than unranked firms. ERCs have long been used to measure the relationship between equity returns and the information contained in a firm's financial reporting. If investors view the level of a firm's commitment and actions in regard to social issues as an indication (or signal) of the trustworthiness of said firms, they would expect that the information contained in reports of firms who are not as invested in these societal issues. Thus, we would expect to find larger ERCs in firms that are ranked by JUST when compared to non-ranked firms. Furthermore, within the population of JUST-ranked firms, we would expect to find larger ERCs for higher ranked firms, we not as compared to lower ranked firms.

The contribution of this paper is two-fold. One, it adds to the literature on the relation between corporate social responsibility and market returns. While previous literature has identified links between equity returns and a firm's commitments to social issues, we use a measure of investor's responsiveness, ERCs, in this context. This extends the

literature in that ERCs can be employed as measure of investors responsiveness both for positive and negative market returns, thus providing insight into the perceived financial reporting quality of these firms. Second, the use of the JUST rankings provides a unique framework within which this relation can be tested. In the past, many of the ESG measures were self-reported by firms. This leaves a gap to fill in the literature as there is no indication if the general public perceived that these firms are socially responsible. In contrast. the JUST rankings are compiled through the polling of everyday Americans to assess which social issues are of importance to them. Thus, the population of those who shaped these rankings is roughly the same as the population of those who invest, providing the opportunity to close this gap in the literature by examining whether people's ideas of socially responsible firms actually translate into greater equity market participation. In short, this paper aims to explore the following research question – do investors value responsiveness to social issues by firms, insofar that it increases their confidence in the quality of earnings announcements of highly social conscious firms.

The remainder of the paper is organized as follows. Section 2 is devoted to a review and discussion of the JUST rankings and the background literature on corporate social responsibility and earnings response coefficients. Section 3 presents the hypotheses development, followed by the research design in section 4. Section 5 provides information on sample selection and summary statistics. Results and discussion and are presented in section 6. Finally, a summary and conclusion are provided in section 7.

2. Background and Literature Review

2.1 JUST Capital Rankings

JUST Capital was founded in 2013. According to their website, its mission is "to build a more just marketplace that better reflects the true priorities of the American people". According to their most recent Form 990 filing, its "definitive polling, rankings, indexes and data empower all market participants - workers, investors, business leaders, consumers, advocacy groups - with the information they need to support, purchase from, invest in, and work for companies that perform best on the issues they care about." To that end, JUST companies are likely to have better pay, create jobs at a higher rate, give more charity and do more to reduce greenhouse gas emissions than non-JUST ranked firms.

JUST Capital ranks firms on a model weighted along seven broad issues including workers (25%), customers (18%), products (14%), environment (13%), jobs (12%), communities (11%) and leadership & shareholders (8%).² Each of these issues are then broken down further into components. In total, there are 36 components. In order to produce their ranking, JUST collects and evaluates data regarding each of these components for each company. An aggregate score is assigned to each company based on the data collected and weights assigned to each issue and component. In total, JUST ranks 890 firms of the Russel 1000, with the first 801 ranked ordinaly from 1-801 and the last 89 grouped together as the "bottom 10%".

2.2 Corporate Social Responsibility and Financial Performance

The relation between a firm's social responsibly actions and its financial performance has been the subject of much debate in both the marketing and accounting literature. Dating all the way back to Friedman (1970), the original literature stated that imposing social expenses on corporations was damaging to the ultimate business goal of increasing profits.³ Early studies argued that companies that incur costs for socially responsible actions were put at a disadvantage relative to firms that did not incur such costs. Arlow and Gannon, writing in 1982, concluded by saying: *"social responsiveness is subordinate to other corporate goals that are primarily economic in nature"*. They continued to assert that *"Although business firms seem committed to socially responsible actions, the research studies do not provide strong support for a positive relationship between social responsiveness and economic performance."* However, others countered with the claim that these costs were minimal and firms actually benefited from the socially responsible actions. Auppele et al. (1985) showed that the results varied depending on the measure of corporate responsiveness. Similarly, later studies such as Jones and Wicks (1999) and McWilliams and Siegel, (2001) concluded that it was in the best interest of firms, even from a profit motive perspective to be socially conscious and responsible.

From a theoretical standpoint, Mackey et al. (2007) develop a model showing that firms that are socially responsible will have higher firm value that those that are not. Gamerschlag et al. (2007) provide evidence displaying shareholders desire for disclosure of CSR information in financial reports. Finally, Alniacik et al., (2011), using a between subject experimental design, find that positive CSR information increases the likelihood that potential investors will invest in the company.

Empirically speaking, McGuire et al. (1988) and Stanwick and Stanwick (1998), among others, show that a firms' financial performance is closely related to perceptions of a firm's social responsibility. Pava and Krausz (1996) show that firms which have been perceived as having met social-responsibility criteria have generally been shown to have financial performance at least on par with, if not better than, other firms. Sasyetki (2015) show that strategic CSR which is disclosed in financial statements have a positive effect on financial performance. In a recent experimental study, Shen et al. (2016) evaluated the influence of CSR disclosure assurance on the investment decisions of nonprofessional investors in China and found that CSR reports that have been assured by an external third party increases the willingness of nonprofessional investors to invest in the firm.

There are a number of later metadata analyses that overwhelmingly conclude that it is beneficial from a profit standpoint for firms to be socially responsible. Margolis and Walsh (2001) look at 60 different studies and found that 53% of these studies documented a positive relationship, 24% found non-significant effect, 19% produced mixed results and only 5% showed a negative effect. Using 52 studies on this topic, Orlitzky et al. (2003) show that it is generally the case that corporate social performance is positively related to corporate financial performance. Even more recently, Friede et al. (2015) combined findings of over 60 review studies on the topic which amounted to more than 2,200 individual studies and concluded that roughly 90% of the studies find a nonnegative relation between ESG and corporate financial reporting with the large majority of studies finding a positive relation.

2.3 Earnings Response Coefficients

Earnings Response Coefficients measure the sensitivity of market returns to earnings surprises. The study of ERCs dates back to seminal works of Ball and Brown (1968) and Beaver (1968). Brown (1993) and O'Brien (1988) show that the most efficient method to verify if expected earnings reflect market expectations is by looking at ERCs.

There is a rich literature indicating how ERCs are related to investors' previous expectations of a firms upcoming earnings report. Imhoff and Lobo (1992) show that ERCs will vary with the amount of ex-ante uncertainty regarding the firm's future prospects. Eason and Zmijewski (1989) report that ERCs are positively associated with revision coefficients (coefficients relating current earnings to future earnings) and negatively associated with expected rates of return. ERCs have also been showed to be influenced by the differential quality of preannouncement information (Verrecchia, 1980), firms size (Atiase, 1985; Freeman, 1987), the effect of earnings predictability (Lipe, 1990), the firm's auditors (Balsam et al., 2003; Teoh and Wong, 1993) and even which exchange the firm's stock trades on (Grant, 1980; Aitase, 1987). Further, the literature has shown that when there is a lack of credibility in regards to a firm's earnings announcements such as in the presence of a qualified earnings report (Choi and Jeter, 1992) or a restatement (Anderson and Yohn, 2002) there is a negative effect on ERCs.

3. Hypotheses Development

If corporate social responsibility is indeed an indication for the trustworthiness of a firm, we would expect that to translate into the perceived quality of financial reporting as well. Equity market participants who receive earnings reports should be more likely to value the information and quality of the earnings reports from firms that are more trustworthy.

The JUST rankings were obtained by surveying the American public on a regular basis to understand what issues represent just corporate behavior, how these issues should be defined, and the relative importance of each (JUST, 2018). JUST then collected and analyzed data on these issues for all the firms in the Russell 1000. Using this data and the relative weights for all of these issues, rankings were developed for each of these firms.

Since the social issues and their weights are objectively determined by surveying the public, the JUST rankings should be a clear indicator of what the public cares about. By extension, companies which score higher in these metrics should be more respected and trusted by the public. This in turn should lead to a better response to an earnings announcement by firms that are JUST-ranked. Thus, we should expect firms that are ranked by JUST to have than more positive response to their earnings report than do non-ranked firms.

This leads to the first hypothesis (stated in the alternative form):

H1: JUST ranked firms have larger earning response coefficients than non-JUST ranked firms.

Although this test is valuable if wishing to examine the difference between firms that are ranked by JUST and those that are not, this does not indicate anything about the rankings themselves. The JUST rankings are based on an aggregate score, by industry, using the 7 issues and 26 components that make up the JUST rankings. Thus, a firm that is ranked higher should be perceived as being more trustworthy that a similar firm that is lower ranked. This

being the case, we would expect that the ERCs for higher ranked firms be higher than those for lower ranked firms. This leads to the second hypothesis (stated in the alternative form):

H2: Firms that are ranked higher by JUST have relatively larger earning response coefficients than lower ranked firms

Additionally, it would be interesting to see to what extent these rankings matter. In other words, is there difference between being ranked in the top 10% vs the top 50% or the top 90%.

4. Research Design

To test *H1*, that JUST- ranked companies enjoy higher ERCs than do non JUST-ranked firms, we estimate the following regression for each firm, by quarter:

$$CAR_{3j,t} = \alpha + \beta_1 UE_{j,t} + \beta_2 JUST_{j,t} + \beta_3 JUST_{j,t} * UE_{j,t} + \beta_4 SIZE_{j,t} + \beta_5 MTB_{j,t} + \beta_6 LEV_{j,t} + \beta_7 LOSS_{j,t} + \beta_8 Q4_{j,t} + \varepsilon$$

$$(1)$$

where $CAR_3_{j,t}$, is the cumulative abnormal return measured over a three-day period surrounding the earnings announcement date for firm *j* in quarter *t*. Following previous literature (Wilson, 2008; Collins & Kothari, 1989), we use a three-day time horizon to measure the market response to earnings surprise.⁴

The first independent variable, UE, is a measure of the firm's quarterly earnings surprise or unexpected earnings. This is measured as the difference between the actual quarterly earnings and the consensus analyst earnings forecast scaled by price. Following previous literature (Chi & Shanthikumar, 2017; Wilson, 2008; Livnat & Mendenhall, 2006), the consensus analyst earnings forecast is defined as the median of analysts' most recent forecasts over 60 trading days before the announcement date. Consistent with prior literature, since the market reacts to this surprise, we would expect β_I to be positive.

JUST is a dummy variable which takes the value of 1 if the firm is a JUST – ranked firm and zero otherwise. Thus, the variable of interest is $JUST_{j,t} * UE_{j,t}$, an interaction variable between this indicator variable and the earnings surprise. The coefficient on this interaction variable will show the incremental response to earnings announcements for JUST-ranked firms. If the market views the information content of earnings for JUST-ranked firms more favorably that for non-JUST ranked firms, the coefficient, β_3 , should be positive.

We also include a number of control variables in the model to control for firm-specific factors that may impact the market return. We include controls for the firm's size (*SIZE*), market to book ratio, (*MTB*), and leverage (*LEV*). *SIZE* is defined as the log of the firm's market value of equity. This study does not make a prediction for the direction of the coefficient on *SIZE* as there is mixed evidence in regards to the relation of a firm's size with the market reaction. Some, such as Chaney and Jeter (1992) find that size is positively related to the market reaction. Others, including Atiase (1985) show that size is negatively related to earnings announcements as these announcements tend to be relatively more informative for smaller firms. As for the other control variables, Collins and Kothari (1989) show that a firm's growth is positively related to the market reaction to a firm's earnings while their financial risk is negatively associated with the market's reaction, while Martikainen (1997) and Dhaliwal et al. (1991) argue that this is not always the case. Thus, we include these variables in the regression model but make no prediction in regards to the sign on either of these variables.

Finally, we include two indicator variable in the regression equation which have been shown to affect market reaction to earnings announcements. The first, *LOSS*, takes the value of 1 if the firm reported a net loss for that quarter, and zero otherwise. The information content of earnings for loss quarters has been shown to be lower than for positive-earnings quarters (Hayn, 1995). Thus, we expect β_7 , the coefficient on *LOSS* to be negatively related to the market reaction. The second indicator variable, *Q4* is set to 1 if the quarter being measure is the fourth quarter of the fiscal year and zero otherwise. Salamon and Stober (1994), show that fourth quarter earnings announcements contain less information content, while Lee et al. (2016) show the opposite to be true. Thus, we do not make a prediction in regards to the sign of the coefficient on *Q4*.

Since there is a debate in the literature regarding the best time horizon to measure market reactions to earnings announcements, we also re-estimate the regression using a five-day time window (see for example Ma et al., 2009 and Ramiah, 2010) by estimating:

$$CAR_{5_{j,t}} = \alpha + \beta_1 UE_{j,t} + \beta_2 JUST_{j,t} + \beta_3 JUST_{j,t} * UE_{j,t} + \beta_4 SIZE_{j,t} + \beta_5 MTB_{j,t} + \beta_6 LEV_{j,t} + \beta_7 LOSS_{j,t} + \beta_8 Q4_{j,t} + \varepsilon ,$$
(2)

where the dependent variable $CAR_{5_{j,t}}$, is the cumulative abnormal return measured over a five-day period surrounding the earnings announcement date for firm j in quarter t.

To test *H2*, that firms that are higher ranked by JUST- capital have larger ERCs than do lower ranked firms, we first begin by dividing the sample in half and estimate the following regression:

$$CAR_{3_{j,t}} = \alpha + \beta_1 UE_{j,t} + \beta_2 JUST_T OP_{50_{j,t}} + \beta_3 JUST_T OP_{50_{j,t}} * UE_{j,t} + \beta_4 SIZE_{j,t} + \beta_5 MTB_{j,t} + \beta_6 LEV_{j,t} + \beta_7 LOSS_{j,t} + \beta_8 Q4_{j,t} + \varepsilon ,$$
(3)

where $JUST_TOP_50_{j,t}$ is an indicator variable which takes on the value of one for firms that are ranked in the top 50% of the JUST- rankings and zero otherwise. We expect the coefficient on $JUST_TOP_50$ to be positive and significant. As before, we re-estimate the regression using $CAR_5_{j,t}$ in place of $CAR_3_{j,t}$.

To further pinpoint the effect of being higher ranked, we then take only the top 10% of JUST-ranked firms, and estimate the following regression:

$$CAR_{3_{j,t}} = \alpha + \beta_1 UE_{j,t} + \beta_2 JUST_T OP_{10_{j,t}} + \beta_3 JUST_T OP_{10_{j,t}} * UE_{j,t} + \beta_4 SIZE_{j,t} + \beta_5 MTB_{j,t} + \beta_6 LEV_{i,t} + \beta_7 LOSS_{i,t} + \beta_8 Q4_{i,t} + \varepsilon$$
(4)

Where $JUST_TOP_10_{j,t}$ is an indicator variable which takes on the value of one for firms that are ranked in the top 10% of the JUST- rankings and zero otherwise. Here too, we expect that the coefficient on this variable will be positive and significant. Again, we re-estimate the regression using $CAR_5_{j,t}$ in place of $CAR_3_{j,t}$.

Finally, to test how far reaching the effect is and if and at what level of ranking the effect wears off, we divide the sample of JUST-ranked firms into deciles and estimate the following regression:

$$CAR_{3j,t} = \alpha + \beta_1 UE_{j,t} + \sum_{N=2}^{10} \beta_N JUST_DECILE_N_{j,t} + \sum_{N=11}^{19} \beta_N JUST_DECILE_N_{j,t} * UE_{j,t} + \beta_{20}SIZE_{j,t} + \beta_{21}MTB_{j,t} + \beta_{22} LEV_{j,t} + \beta_{23} LOSS_{j,t} + \beta_{24}Q4_{j,t} + \varepsilon$$
(5)

where *JUST_DECILE_N* refers to the firms in each of the 10 deciles. We omit Decile 1 (the top 10%) as this will be the benchmark and will be captured in β_1 , the coefficient on *UE*. Here, we expect that the coefficient on each D_N^*UE will be greater than the coefficient on D_{N+1}^*UE . Additionally, we conduct F-tests to test the difference between the coefficient on each D_N^*UE and the coefficient on D_1^*UE as well as F-tests to test the difference in coefficients of each D_N^*UE and the coefficient on the next decile, D_{N+1}^*UE . Finally, we re-run these tests using *CAR_5_{j,t}* in place of *CAR_3_{j,t}*.

5. Data Selection and Summary Statistics

5.1 Data Selection

The JUST Capital rankings are publicly available on their website.⁵ Market return data is obtained from CRSP, analyst forecast data from IBES and financial data from COMPUSTAT. Since the JUST data is for the year 2018, this study examines the previous 10 years (beginning in 2008) as well as 2018. After merging data from all the databases, and removing observations with insufficient data, as well as firms from highly regulated industries such as banking, insurance and utilities, the final sample consists of 101,649 firm quarter observations. Of these, 25,318 quarter observations are from 854 JUST-ranked companies and 75,331 are from non-ranked firms. For the 11 years that the sample covers, the percentage of JUST rank firms range from 20.43% to 28.89% with the mean being 25%.

5.2 Summary Statistics

Descriptive statistics for all variables are presented in Table 1. As expected, JUST firms, which are made up of the Russell 1000 are larger and have a greater market to book value. They also have less instances of reporting a loss than do non-JUST firms. Interestingly, the mean and median cumulative abnormal returns (both for the three- day and the five- day window) are positive for JUST firms and negative for non-JUST firms. This would seem to indicate that earnings announcements from larger companies create stronger market reactions than those of smaller firms.

			Non JU	U ST firms				
N = 76,331	CAR_3	CAR_5	UE	SIZE	MTB	LEV	LOSS	Q4
Mean	-0.0003	-0.0004	-0.0034	6.7598	2.7259	0.5488	0.2506	0.2163
Median	-0.0004	-0.0007	0.0005	6.7463	1.7443	0.5404	0.0000	0.0000
SD	0.0998	0.1120	0.0416	1.5400	4.8571	0.2674	0.4334	0.4117
P10	-0.1017	-0.1146	-0.0119	4.8209	0.6384	0.1918	0.0000	0.0000
P90	0.1014	0.1131	0.0105	8.6339	5.7721	0.8974	1.0000	1.0000
JUST firms								
$\frac{0001111100}{N = 25318}$	CAR 3	CAR 5	UE	SIZE	MTR	LEV	LOSS	04
Mean	0.0045	0.0051	0.0006	9 1065	4 2396	0.5886	0 0774	0 2319
Median	0.0036	0.0050	0.0005	9.0004	2 9048	0.5854	0.0000	0.0000
SD	0.0677	0.0725	0.0108	1 2219	6 0752	0.2202	0.2672	0.4221
P10	-0.0688	-0.0736	-0.0011	7.6805	1.0790	0.3057	0.0000	0.0000
P90	0.0804	0.0854	0.0036	10.8147	9.1199	0.8686	0.0000	1.0000
All firms								
N = 101,649	CAR_3	CAR_5	UE	SIZE	MTB	LEV	LOSS	Q4
Mean	0.0009	0.0010	-0.0024	7.3443	3.1029	0.5587	0.2075	0.2202
Median	0.0006	0.0008	0.0005	7.2797	1.9909	0.5549	0.0000	0.0000
SD	0.0929	0.1036	0.0365	1.7841	5.2284	0.2570	0.4055	0.4144
P10	-0.0931	-0.1035	-0.0081	5.0783	0.7027	0.2132	0.0000	0.0000
P90	0.0951	0.1054	0.0083	9.7463	6.7081	0.8935	1.0000	1.0000
Table 2. Panel B:	Test of differ	ence in mear	18					
		Non J	IUST firms	JUST firm	ıs Di <u>f</u>	ference		
	CAR	3 -	0.0003	0.0045	0.004	48 ***		
	CAR	5 -	0.0004	0.0051	0.005	55 ***		
	UE	-	0.0034	0.0006	0.004	40 ***		
	SIZI	Ε (5.7598	9.1065	2.340	57 ***		
	MTB 2.7259		4.2396	1.513	37 ***			
	LEV	/ ().5488	0.5886	0.039	98 ***		
	LOS	S ().2506	0.0774	-0.17	32 ***		
	Q4	().2163	0.2319	0.01	56 ***		
	Ν	~	76,331	25,318				
	*** *	*. * indicate	es significance	e at the 1%. 5%	%. 10% leve	.1		

Table 1. Panel A: Descriptive Statistics

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Correlations for all variables are presented in Table 2. As would be expected from the previous literature, unexpected earnings (UE) is positively correlated with both CAR_3 and CAR_5. Consistent with Atiase (1985) and Hayn (1985), respectively, we observe that SIZE and LOSS are both negatively correlated with UE. This is because the level of surprise in earnings announcements are generally of a lower magnitude for larger firms and for firms that report a loss.

	CAR3	CAR5	UE	MTB	SIZE	LEV	LOSS
CAD5	0.0012						
CAKS	0.9012						
	0.0000						
UE	0.1342	0.1275					
	0.0000	0.0000					
SIZE	-0.0038	-0.0062	0 1114				
	0.2243	0.0478	0.0000				
MTD	0.0085	0 0088	0.0208	0 1462			
MIID	-0.0085	-0.0088	0.0208	0.1402			
	0.0068	0.0052	0.0000	0.0000			
LEV	0.0065	0.0075	-0.0554	-0.0291	0.0737		
	0.0380	0.0162	0.0000	0.0000	0.0000		
LOSS	-0.0948	-0.0887	-0.2498	0.0245	-0.3722	-0.0380	
	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	
Q4	0.0130	0.0163	-0.0094	-0.0007	0.0294	0.0085	-0.0156
	0.0000	0.0000	0.0028	0.8240	0.0000	0.0068	0.0000

Table 3. Correlations (coefficients and p-values)

6. Results and Discussion

To test H1, we estimate a regression for the entire sample with an indicator variable *JUST* as well as an interaction variable between *JUST* and *UE*. Results from this regression are presented in Table 3. The coefficient on *JUST*UE* when using *CAR_3* (column 1) is positive with a very large t-value of 35.89. Similarly, when testing using a five-day window (column 2), we again find a large significant (t=34.32) positive coefficient. This is consistent with *H1*, namely, that JUST-ranked firms have larger ERCs than do non-JUST ranked firms.

	(1)	(2)	
	CAR_3	CAR_5	
UE	0.292***	0.313***	
	(35.89)	(34.32)	
JUST	0.003***	0.005***	
	(3.60)	(4.94)	
	***	***	
JUST_UE	6.559	6.642	
	(27.11)	(24.56)	
SI7F	0.004***	0.004***	
SILL	(17.18)	(17.66)	
	(-17.16)	(-17.00)	
MTB	-0.000	-0.000	
	(-0.03)	(-0.21)	
LEV	0.004^{***}		
	(3.57)	(3.86)	
	0.010***	0.010***	
LOSS	-0.018	-0.019	
	(-23.27)	(-21.96)	
04	0.003****	0.005***	
~	(4.86)	(5.89)	
	. /		
Intercept	0.027***	0.030***	
	(16.47)	(16.50)	
 <i>t</i> statistics in pa	(16.47) arentheses * $p < 0.10$, *	(16.50)	

Next, to test *H2*, we divide the JUST-ranked firms in half and estimate a regression with the variable $JUST_TOP_50$ as well as the interaction variable $JUST_TOP_50 * UE$. Results from this regression are presented in Table 4. Here, again we find a positive coefficient on $JUST_TOP_50 * UE$, significant at the 5% level, both when the return window is three days (t=2.47) and five days (t=2.06). This confirms *H2*, that the higher a firm is ranked, the larger its ERCs will be.

Table 5. Regression analysis JUST TOP 50

	(1)	(2)	
	(I)	(2)	
	0.505***	0.445***	
02	(10.90)	(8.95)	
JUST_TOP_50	0.001	0.002	
	(0.96)	(1.61)	
IUST TOP 50*UE	0.219**	0 196**	
	(2.47)	(2.06)	
	***	***	
SIZE	-0.004	-0.005	
	(-10.06)	(-11.32)	
MTB	0.000	0.000	
	(1.21)	(0.71)	
LEV	-0.002	-0.000	
	(-0.79)	(-0.13)	
LOSS	-0.007***	-0.007***	
	(-4.42)	(-3.95)	
04	0.001	0.002**	
Q4	0.001	0.002	
	(1.47)	(2.31)	
Intercept	0.040^{***}	0.047***	
	(11.31)	(12.36)	

 $\begin{aligned} CAR_{j,t} &= \beta_0 + \beta_1 UE_{j,t} + \beta_2 JUST_TOP_50_{j,t} + \beta_3 JUST_TOP_50_{j,t} * UE_{j,t} + \beta_4 SIZE_{j,t} + \beta_5 MTB_{j,t} + \beta_6 LEV_{j,t} \\ &+ \beta_7 LOSS_{j,t} + \beta_8 Q4_{j,t} + \varepsilon \end{aligned}$

t statistics in parentheses ${}^{*}p < 0.10$, ${}^{**}p < 0.05$, ${}^{***}p < 0.01$

To further narrow down the effect, we repeat the test with the variable $JUST_TOP_10 * UE$, measuring the interaction of the unexpected earnings with only the top 10% of JUST- ranked firms. If a higher ranking is indeed related to larger ERCs, we would expect to have even larger coefficients for this variable than we found for the top 50%. Table 5 presents the results from this regression. We find positive coefficients of 1.295 (t=5.66) for the three-day window and 1.66 (t=6.77) for the five-day window. When compared to the previous test of the top 50%, these coefficients are much larger (1.295 vs .219 and 1.65 vs .196) as well as more significant (5.66 vs 2.47 and 6.77 vs 2.06).

Table 6. Regression analysis JUST TOP 10

	(1)	(2)	
	CAR_3	CAR_5	
UE	0.524***	0.447***	
	(12.96)	(10.32)	
JUST_TOP_10	0.002	0.002	
	(1.56)	(1.49)	
JUST_TOP_10*UE	1.295***	1.660***	
	(5.66)	(6.77)	
SIZE	-0.004***	-0.005***	
	(-10.64)	(-11.69)	
МТВ	0.000	0.000	
	(1.19)	(0.70)	
LEV	-0.001	-0.000	
	(-0.69)	(-0.05)	
LOSS	-0.007***	-0.007***	
	(-4.46)	(-3.96)	
04	0.001	0.002**	
~	(1.45)	(2.29)	
Intercent	0.042***	0.048***	
Intercept	(11.64)	(12.53)	

t statistics in parentheses p < 0.10, p < 0.05, p < 0.01

Having established that higher ranked firms have larger ERCs than lower ranked firms and that this effect is stronger in the highest 10% of firms than it is in the top 50%, we now examine at which point this effect weakens. We divide the JUST firms into 10 deciles and estimate a regression including all 9 of the 10 deciles. The top decile, that is the top 10%, will be accounted for in the coefficient on UE. Results from this regression are presented in Table 6. Panel A presents the results when the dependent variable is the 3-day cumulative return. Column 1 presents the coefficient for each interaction between the unexpected returns and the decile. Column 2 is the interpretation of the coefficient; that is the sum of the coefficient on UE (the top 10% of firms) alone plus the incremental of the interaction.

We also test the coefficient of each decile vs the coefficient on the 10% decile. Significance of these tests are presented in Column 3. We find that for each decile the difference coefficient from the coefficient of 10% decile is significant at the 1% level. This is consistent with the earlier results, namely that firms ranked in the top 10% have significantly larger ERCs than do lower ranked firms.

Finally, we perform an F test for the coefficient on each decile *N*, against the coefficient on the next decile, N+1 to see where the differences fade out. Significant levels from these results are presented in column 4 of Table 6. Interestingly, we find that while the difference between the Decile 1 (the top 10%) and Decile 2 (11%-20%) are significant at the 1% level, the difference between Deciles 2 and 3 and Deciles 3 and 4 do not produce a significant F-statistic. This indicates that while being ranked in the top 10% is related to larger ERCs, there is not much difference between being in the 11% to 40% range. However, between Deciles 4 and 5 there again is a significant F-value while between deciles 5 and 6 there is not. This would indicate that there is still value of being in the top 40% vs being below the 60^{th} percent.

Lastly, we observe that the F-statistic on the test between Deciles 6 and 7 is significant at the 1% level, while the rest of the F-statistics are small and either significant only at the 10% level or insignificant, indicating that being ranked below the top 60% has a much bigger relation to the ERCs than being ranked in the bottom 40%.

In sum, these results would indicate that the effect that JUST rankings have on ERCs can be divided up into four segments. The top 10% has by far the largest effect. After that, companies ranked within the 40th percentile would be the second most effected. The third effect is from the 41st percentile until the 60th. Below the 60th has very little effect, although those ERCs still tend to be bigger than the non-JUST ranked firms. Clearly, the benefit of being ranked by JUST increases as firms rise in the rankings but the increase in benefits is not linear.

Table 6, Panel B presents results of this regression when using CAR_5 as the dependent variable. In general, the findings are similar to what was presented in Panel. A. Interestingly, here, there is still a difference between Decile 6 and 7 at the 5% level as well as between Decile 9 and Decile 10 (bottom 10%). This would indicate that for five-day trading windows, there is still some effect for firm ranked below the 60th percentile all the way down to the 90th percentile.

	(1)	(2)	(3)	(4)
			$\beta_{n=1}$ vs β_n	β_n and β_{n+1}
	CAR_3		F stat	F stat
UE	1.738***			p < .01
	(12.04)			
JUST_D2*UE	-1.049***	0.689***	p < .01	
	(-5.23)	(4.93)		
JUST_D3*UE	-1.068***	0.670^{***}	p < .01	
	(-4.15)	(3.14)		
JUST_D4*UE	-0.661***	1.078***	p < .01	p < .01
	(-3.00)	(6.48)		
JUST_D5*UE	-1.719***	0.019	p < .01	
	(-8.29)	(0.13)		
JUST_D6*UE	-1.606***	0.132*	p < .01	p < .01
	(-10.08)	(1.94)		
JUST_D7*UE	-1.099***	0.640***	p < .01	p < .10
	(-5.95)	(5.54)		
JUST_D8*UE	-0.786***	0.952***	p < .01	
	(-3.90)	(6.76)		
JUST_D9*UE	-1.017***	0.721***	p < .01	
	(-5.16)	(5.35)	-	
JUST_D10*UE	-1.113***	0.625***	p < .01	
_	(-5.56)	(4.49)	•	
Intercent	0.039***			
1	(10.11)			

Table 7. Panel A: Regression analysis by JUST Decile using CAR_3

$CAR_5_{j,t} = \alpha +$	$\beta_1 U E_{j,t} + \sum_{n=1}^{10} \beta_n J U S T$	$T_DECILE_N_{j,t} + \sum_{j=1}^{19} \mu_j$	3 _N JUST_DECILE_N _{j,t} *	$UE_{j,t} + \beta_{20}SIZE_{j,t}$
	$\frac{\sum_{N=2}}{+ \beta_{21}MTB_{j,t} + \beta_{22} LE}$	$\frac{\sum_{N=11}^{N=11}}{V_{j,t} + \beta_{23} LOSS_{j,t} + \beta_{24}0}$	$Q4_{j,t} + \varepsilon$	
	(1)	(2)	(3)	(4)
			$\beta_{n=1}$ vs β_n	β_n and β_{n+1}
	CAR_5		F stat	F stat
UE	1.939***			p < .01
	(12.53)			
JUST D2*UE	-1.246***	0.694***	p < .01	
	(-5.79)	(4.63)	I	
JUST_D3*UE	-1.206***	0.733***	p < .01	
	(-4.37)	(3.21)		
	1.072***	0.00	. 01	. 01
JUSI_D4*UE	-1.053	0.886	p < .01	p < .01
	(-4.40)	(4.97)		
JUST_D5*UE	-2.249***	-0.310*	p < .01	p < .05
	(-10.12)	(-1.94)		
JUST_D6*UE	-1.880***	0.059	p < .01	p < .01
	(-11.01)	(0.80)		
IIIST D7*I/F	-1 396***	0 543***	n < 01	n < 05
J 051_D/ 0L	(-7.06)	(4.39)	p < .01	p < .05
	(1.00)	(1.57)		
JUST_D8*UE	-0.982***	0.957***	p < .01	
	(-4.55)	(6.33)		
JUST_D9*UE	-1.151***	0.788***	p < .01	p < .05
	(-5.45)	(5.46)		
JUST D10*UE	-1.578***	0.361**	p < .01	
	(-7.36)	(2.42)	r	
	× /	× /		
Intercept	0.046***			
	(11.12)			

Table 8. Panel B: Regression analysis by JUST Decile using CAR_5

t statistics in parentheses * p < 0.10, ** p < 0.05, *** p < 0.01

7. Conclusion

Although previous studies have shown that investors value firms that invest in corporate social responsibility, this study measures this relation using earnings response coefficients. Thus, it examines the public's confidence in the reliability of earnings reports for firms that are reputed to be socially conscious. Moreover, the metric used to measure a firms' social responsiveness, the JUST- rankings is a relatively new measure which has the distinction of being directly linked with social issues that are most valued by the public at large, creating a unique setting to examine the relation between CSR and market returns.

Based on the findings in this study, it does indeed seem to be the case that investors value companies that score higher on these measures of social responsibility. This is especially true of the top 10% of firms as well as the top 50%. However, the effect seems to weaken somewhat for firms that are no longer in the top 40% of the rankings. We conclude from this study that investors do take into account the level of a corporation's social responsibility when evaluating earnings announcements. Investors perceive increased quality and decision usefulness in earnings announcements for firms that are more socially responsible.

While this paper adds to the ongoing study of the link between CSR and market performance, there are still many avenues for research in this area. Perhaps one such avenue would be to examine the individual components that make up the JUST rankings and see how the market reacts to firms that score higher (or lower) on one or more particular components. Another possible area of research would be a look-forward approach to see if and how JUST-rankings affect investor behavior moving forward.

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Notes

Variable Name	Variable Description
JUST	An indicator variable equal to 1 for firms that are ranked by JUST Capital and 0 otherwise.
CAR_3 (CAR_5)	Cumulative abnormal return surrounding the earnings announcement for firm <i>j</i> at quarter <i>t</i> , measured over a three-day (five-day) window, where the abnormal return is the firm's return less the CRSP value-weighted market return.
UE	The quarterly earnings surprise for firm j at quarter t 's announcement date, scaled by actual earnings, where expected earnings is based on the median of analyst forecasts outstanding within 60 days prior to the day before the earnings announcement.
SIZE	The log of the firm's market value of equity.
MTB	The firm's market to book ratio.
LEV	The firm's leverage ratio.
LOSS	An indicator variable equal to 1 if actual reported earnings per share are less than zero, and 0 otherwise.
<i>Q</i> 4	An indicator variable equal to 1 if the earnings announcement is for the fourth quarter of the firm's fiscal year, and 0 otherwise.

Note 1. Although the universe of JUST ranked companies is composed of firms in the Russell 1000 (roughly equivalent to the 1,000 largest, publicly traded U.S. companies by market capitalization) due to data availability as well as companies that have been acquired, in actuality only 890 companies were ranked by JUST Capital in 2018.

Note 2. These weights, provided by JUST Capital (<u>https://justcapital.com/polling/</u>) add up to 101%, presumably due to rounding.

Note 3. Just recently, The Business Roundtable, an association of the chief executive officers of nearly 200 of America's largest and most influential companies, released a new Statement on the Purpose of a Corporation (available at

<u>https://opportunity.businessroundtable.org/wp-content/uploads/2019/09/BRT-Statement-on-the-Purpose-of-a-Corpor</u> <u>ation-with-Signatures-1.pdf</u>), signed by 181 CEOs in which they essentially reject Friedman's idea. In the statement, they commit to delivering value to customers, investing in employees, dealing fairly and ethically with suppliers and supporting communities. The very last commitment they mention is to generate long-term value for shareholders.

Note 4. When performing an event study of earnings announcements, the day following the announcement day is usually added to the event window to capture the market reaction in the event the announcement occurs after trading hours. Additionally, one day prior to the announcement day is added to capture the market reaction to possible information leakages preceding the official earnings announcement. Adding too many days may lower the accuracy of the test due to the possibility of confounding effects from other market events (MacKinlay, 1997). To examine the sensitivity of the earnings announcement to different lengths, we test both for abnormal returns for a three-day window (-1, +1) and a five-day window (-2, +2).

Note 5. www.justcapital.com