Stock Market Development and Economic Growth: Empirical Evidence From an Institutional Impaired Economy

Maku Affor Owen¹

¹ Delta State University, Abraka, Nigeria

Correspondence: Maku Affor Owen, Delta State University, Abraka, Nigeria. E-mail: makuowen@gmail.com

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Abstract

The research investigated the relationship linking stock market development and economic growth from 1985 to 2018. In measuring growth, Gross domestic product (GDP) was adopted, while stock market was surrogated by turnover ratio, market-capitalization, and value of share- traded, sourced from the Central Bank of Nigeria (CBN) and the Security and Exchange Commission Database. The inclusion of money supply (M3) captured innovation (financial) in the monetary sector. In investigating the aforementioned relationship, the ARDL Bound test methodology was adopted. Empirical results from the investigation confirm the existence of a long-run relationship between stock market development and growth. Similarly, there was a positive relationship between indices of stock market development and growth, albeit statistically insignificant. The study concluded that financial institutions should concentrate on financial innovation in other dimensions in other to boost stock market performance that will result in sustainable growth.

Keywords: financial market, economic growth, stock market, money supply, financial innovation

JEL: N21, O40, D53, E51, O53

1. Introduction

As Africa's financial markets tend towards a higher degree of complexity, the stock markets' relevance cannot be overemphasized. Shittu (2012) posits that stock markets' contribute to growth through certain channels such as creation of liquidity, mobilization of savings for private and public sectors, risk diversification, and improved dissemination of information. Through the delivery of these services, the advancement of economic growth can be enhanced. In most developing economies, the emergence and growth of stock markets have reliably increased over time. Amid stock market size and illiquidity, its continuous turn of events and existence may have an expanding impact for growth. Pardy (1992) argued that even in LDCs, capital markets/institutions can mobilize savings (domestic) and distribute funds more efficiently. In this way, the stock-market can play its principal role in stimulating LDCs growth by attracting and redirecting investment to where it is needed.

Several studies relating to the Nigerian economy have constantly and consistently appraised the responsibility of the stock market on economic growth. A number of studies have observed the potential of the Nigerian stock market to stimulate growth (Olofin & Afangideh, (2008); Ogunmuyiwa, 2010 and Ezeoha, Ogamba & Onyiuke, (2009)). These studies have found positive effects, flowing from the stock market to economic growth. Odhiambo (2009) argues that the development of the stock market Granger-causes economic growth while Ndako (2009) conversely asserts empirically that economic growth granger causes stock market development, with similar evidences from time series data adopted above. In addition, recent studies including (Osakwe & Ananwude, (2017); Okonkwo et. al, (2015) and Adigwe et. al, (2015); Carpenter & Whitelaw, (2017); Babjide et. al, (2016); Brown & Nyeche, (2016); Pan & Mishra, 2018) empirically confirmed that the development of the stock market positive contribution to economic growth. On the basis of the assumption, this study critically analyzed the relationship between stock market development and economic growth in Nigeria. This research paper specifically assesses the stability, depth and efficiency through its interplay with financial innovation on the growth of the Nigerian economy.

The remainder of this paper is subdivided into four different sections. Section two covers relevant analytical and theoretical research, section three discusses the methodology adopted; section four focuses on presenting the results and discussion of findings, while section five provides relevant policy recommendations.

2. Literature Review

The relevance of the contribution of the financial system in improving and promoting development cannot be overemphasized. The financial sector comprises the Central Bank, commercial banks, Investment companies, brokerage firms, discount houses, and the stock exchange, to name just a few. These institutions trade in financial instruments that include stocks, derivatives, foreign and domestic currencies, shares, and so on, and mobilize funds from surplus to units of deficit in the same process. This helps to boost investment and increase production for businesses and corporations, thereby accelerating growth. The debate on the functioning of the financial system began with Schumpeter (1912), who argued that banks operate in a well-functioning financial system to stimulate economic growth by stimulating technological innovation by identifying and financing entrepreneurs with the best opportunity to launch innovative products, in addition to the production process. Levine (1991) was of the opinion that a well-developed stock market possesses the capacity to absorb liquidity shocks and productivity shocks of businesses. In the same vein, Levine & Zervos (1998) and Khan & Senhadji (2000) have pointed out that the establishment of the stock market has made significant contribution to the growth of financial institutions in emerging-market economies. The expansion of the financial sector (including the stock market) is therefore assumed to make substantial contribution to growth.

2.1 Stock Market Performance and the Nigerian Economy

Table 1 and 2 shows key stock market growth metrics (stock market size, depth, and market stability) for the period 1985-2018. Prior to onset of the global financial crisis and financial sector reforms in 2005, the Nigerian stock market was well adjusted, as market indices (market capitalization, market turnover, value of shares traded and the all-share index) rose from low to historically high levels. For instance, market capitalization increased from N5bn in 1985 to N18.3bn in 1990, and further increased to N165.10bn in 1995, representing an increase of about 13.56 percent as a share-of-GDP. In 2000 and 2005, market capitalization was set at N379.71bn and N2, 066.80bn, representing an increase of 51 percent as a share of GDP. This period marks the entrant into the era of post consolidation. The rapid development of the stock market, however, was primarily the result of stock market trading which resulted in just few stocks, accounting for a substantial part of overall market capitalization. There are unsmilingly informational and disclosure deficiencies for other securities, as well as stern flaws in the lucidity of transactions in the market, away from these active-traded shares. Similarly, market turnover resumed a growing trend, increasing from 2000 to 2005 to 6 percent to 10.1 percent. The all-share index also performed well, rising from 117points in 1985 to 3,815 points in 1995 and further to 6,701 points in 2000. The all-share index peaked at 50,424 points in 2008 and was steadily decreasing to 37,186 points in 2018. Regrettably, as market metrics deteriorated quickly, the boom experienced in the market was upturned. For example, market capitalization in 2018 currently stands at N16,185.7billion, indicating a decreasing GDP ratio of 7.36 percent, while market turnover reported a decline of around 7 percent in 2018 compared to 10.10 percent in 2010 (CBN Annual Report, 2019).

| | Ratio to GDP (percent) | | | |
|------|------------------------|-------------------------------------|-----------------------------|-------------------------------------|
| Year | Stock market turnover | Stock market to total value- traded | Stock market capitalization | Stock market capitalization (World) |
| 1995 | 1.07 | 0.18 | 17.65 | 66.04 |
| 2005 | 8.78 | 1.10 | 12.62 | 92.52 |
| 2010 | 10.10 | 1.40 | 13.91 | 88.48 |
| 2015 | 8.17 | 0.82 | 10.10 | 94.96 |
| 2016 | 5.36 | 0.37 | 7.36 | 97.46 |
| 2017 | 5.87 | 0.58 | 9.90 | 111.83 |
| 2018 | 8.20 | 0.65 | 7.93 | 92.92 |

Table 1. Indicators of stock market development for Nigeria

Source: Global Financial Development Indicators, World Bank



Figure 1. Nigeria's market capitalization ratio to World's market capitalization

Source: World Development Report, 2019

| Table 2. Stock marke | t development in | dicators for Ni | igeria (1985 – 2018) |
|-----------------------|------------------|-----------------|----------------------|
| 1 ubie 2. biook marke | t development m | areators for 10 | .genu (1905 2010) |

| Year | Market Capitalization (₦' Billion) | All Share Index | Turnover Ratio (%) | Value of stocks traded (N' Billion) |
|-------------|---------------------------------------|-----------------|-----------------------|--|
| <i>1985</i> | 5.075 | 117.283 | 0.5 | 0.316 |
| 1990 | 18.334 | 423.658 | 0.7 | 0.225 |
| 1995 | 165.099 | 3815.117 | 1.0 | 1.838 |
| 2000 | 379.713 | 6701.175 | 6.0 | 28.153 |
| 2005 | 2066.798 | 22876.717 | 10.1 | 262.935 |
| 2010 | 6493.303 | 24775.512 | 10.1 | 799.911 |
| 2015 | 10521.630 | 30867.195 | 9.7 | 978.047 |
| 2016 | 9151.608 | 26624.077 | 6.2 | 620.018 |
| 2017 | 11150.788 | 32161.113 | 9.3 | 1078.492 |
| 2018 | 13499.618 | 37186.112 | 7.0 | 1,284.976 |

Source: NSE and CBN

Pagano (1993) argued that the stock-market contributes to the mobilization of domestic savings by strengthening the array of financial instruments available to savers in diversifying their portfolios, thereby providing a significant source of investment capital at a relatively low cost. A well-functioning and liquid stock market that provides investment opportunities to diversify unsystematic risks can improve capital's marginal productivity. Taking a retrospect at the World Federation of Exchanges in Paris, it is shown that overall global stock market capitalization, rose from US\$40.4 in 2005 to US\$50.9 trillion in 2010 (World Bank, 2019). As indicated in figure 1, the overall world stock market capitalization saw a steady rise per year from 92.52 percent in 2005 to 111.83 percent in 2017 and decreased to 92.92 percent in 2018 in terms of its share-of-GDP in terms of its share in GDP. Total global stock market capitalization, which in 1995 amounted to US\$ 66.04 trillion; increased to US\$ 111.83 trillion in 2017, nearly double the rise over seventeen (17) years. By comparison, stock market capitalization in Nigeria is experiencing an increasing trend. Successive periods (2016-2018) for market capitalization and market turnover ratio experienced undulating patterns which signifies poor performance underdevelopment of the stock market. The rapid development of global stock indicators (that is, market capitalization) has attracted serious attention of policymakers; thus focusing on the reasons for the under performance of the stock market in developing countries around the world over the last few decades, giving room to explore other dimensions that will result in the effective development the stock market and growth of developing economies (Deb & Mukherjee, 2008)

2.2 Empirical Review

Studies have produced mixed results with respect to country specific and cross-country research. Studies including Ikikii & Nzomoi (2013), Bernard and Austin (2012), Ojo & Adeusi (2012), Edame & Okoro (2013), Kolapo & Adaramola (2012), Afolabi (2015), Wang & Ajit (2013), Okoye & Nwisienyi (2013), Jibril et. al (2015), Echekoba, Ezu & Egbunike (2013), Haque (2013), Fynn (2012), Sinha, Joshi, Venkartaraman, Padmanabhna & Ravi (2015), Ologunwa & Sadibo (2016), Odo, Anoke, Onyeisi & Chukwu (2017), Karim & Chaudhary (2017) either found that the development of stock markets had a positive impact on economic growth or had no significant effect on growth.

S/N Author Country(ies) Studied Period **Estimation Technique** Findings **Country Specific Review** China (Hong-Kong) 1 Ho & Odhiambo 1981-2010 ARDL bounds test There was a uni-directional causality flowing from market (2015) capitalization to growth, and a causal flow from growth to stock market-turnover in the long and short-run and a casual flow from stock market turnover to growth in the short term 2 Bernard & Austin Nigeria 1994-2008 OLS Market capitalization and market liquidity adversely (2012)affect growth 3 Adeusi Market capitalization has a effect impact on growth as Ojo & Nigeria Johansen cointegration (2012) test value traded, volume of transactions and all share index negatively affects economic growth 4 Ovat (2012) Nigeria OLS The result showed that stock market development through contributes to Nigeria's growth market-liquidity-based indicators; turnover ratio and total value of shares traded. 5 Wang & Ajit China Quarterly OLS The result showed that development of the stock market (2013)does not contribute positively to growth in developing data (1996-2011) countries, given that the stock-market is mainly administratively market-driven Ikikii & Nzomoi 6 Quarterly OLS Stock market trading and capitalization levels have a Kenva (2013)positive impact on economic growth. time-series data (2000-mid 2011) 7 Marinkovic et al. Serbia 2002-2011 Granger causality Stock turnover ratio to GDP Granger causes real GDP (2013)growth 8 Tang (2013) 1960-2008 Cointegration There was a uni-directional causal relationship flowing Australia and Granger causality Tests from stock prices to growth 9 Edame & Okoro 1970-2010 OLS Studies have shown that the stock market contributes Nigeria (2013)significantly to Nigeria's economic growth. 10 Echekoba et. Nigeria 1999-2011 Multivariate regression The results showed that total market capitalization and all al(2013) analysis share index have a positive impact on GDP-provoked economic growth. 11 Okoye & Nigeria 2000-2010 Multiple regression and The result showed that a significant relationship on GDP Nwisienyi (2013) OLS between all share index, market value and market capitalisation. 12 B.O Zimbabwe 1990-2010 Toda The results showed bi-directional relationship between Ishioro and Yamamoto (2013)Granger non-causality shifts in the stock market and economic growth. technique 13 Afolabi (2015) 1992-2011 Nigeria Multiple regression The result revealed that, during the timeframe under analysis study, the stock market had marginal effect on the economy 14 Jibril et. al (2015) Nigeria 1999-2010 OLS The study found that market capitalization and value trade ratio is negative correlated with economic growth, whereas the turnover ratio has a strong positive

Table 3. Summary of relevant literature

| | | | | | relationship with economic growth. |
|-------|--------------------------------|---|-------------|--|--|
| 15 | Altarturi & Abduh (2016) | Malaysia | - | Granger-causality | The results revealed that there is a bidirectional relationship between Islamic stock markets and Malaysian growth, and appears that the contribution to growth is indirect in relation to its effect on investment. |
| Cross | -Country Review | | | | |
| 16 | Osamwonyi & Kasimu (2013) | Nigeria, Kenya and Ghana | 1989 – 2009 | Johansen co-integration/ Granger causality | In Ghana and Nigeria, no causal link was identified between the development of the stock market and economic growth. In comparison, Kenya had a bi-directional causal link between stock market development and growth. |
| 17 | Haque (2013) | Sri-Lanka Bangladesh, India and Pakistan | 1980-2005 | DOLS | The Stocks market had no impact on per capita GDP growth |
| 18 | Kagochi et al. (2013) | Sub-Saharan Africa | 1991-2007 | Panel causality test | The causal relationship between the development of the stock market and growth has bas shown to be bidirectional. |
| 19 | Hailemariam & Guotai (2014) | 17 emerging and 10 developed countries | 2000-2011 | GMM dynamic panel | Key findings revealed that there is statistically significance between the development of stock markets and economic growth and an indirect improvement of investment behaviour. |
| 20 | Sinha et. al (2015) | 14 selected Asia-pacific countries | 1993-2009 | | The findings showed no major correlations between the returns on a country's stock market and its performance in GDP. However, the findings indicate that in order to achieve a healthy and sustainable quality of life, economic growth in the LDCs must be accompanied by the subsequent improvement in income distribution, social security openness and accountability |
| 21 | Jalloh M. (2015) | Selected African countries | - | Dynamic Panel estimation | Study findings showed that hiking stock market capitalization by a median average of 10% triggers growth in countries surveyed by 5.4%. |
| 22 | Ruwaydah & Ushad (2015) | 9 developing counties in SADC region | 1980-2011 | Pooled Panel | Results indicate a clear correlation existing between stock market development and growth. |
| 23 | Bilal et. al (2016) | 20 Lower-middle income countries | 1990-2012 | Panel estimation | Overall results suggest that stock market development has a major effect on growth. The research also assessed the precise effect on growth of stock market development on growth by controlling variables including financial depth, Investment, foreign direct investment, openness and inflation. |
| 24 | Azam et. al (2016) | Four Asian States (Bangladesh, India, China and Singapore) | 1991-2012 | ARDL Bounds test | The results have shown that economic growth, foreign direct investment, stock market development and inflation are co-iintegrated in the long run |
| 25 | Okoro G.E (2016) | Nigeria and Mauritius | 2006-2010 | Paired-samples T-test statistical method | Empirical results indicate that Mauritanian stock market output was higher than Nigeria's and the same for GDP. However, stock market performance has had a negative impact on economic growth in Mauritius and Nigeria due to the fact that emerging markets are giving attention to the money market while relegating the stock market to the background. |
| 26 | Karim et. al (2017) | 2 In Asian regions (South and East Asia) | 1996-2015 | Linear Panel method | Empirical results indicated that the development of the stock market contributes to some degree to economic growth of the South Asian region, but its effect on East Asian region has proven to be negligible. |

Source: Compiled by Author

3. Methodology

The study adopts the ARDL model fundamentally centered on the following premise. First, regardless of the sample size, the ARDL model is advanced (Ghatak & Siddiki, 2001). Secondly, particularly when the variables are integrated of different order (that is integrated at order zero or one), the method is further suitable. Third, modelling ARDL with the appropriate lags is standard for both the indigeneity problem and serial correlation, and fourth, the model can estimate the short and long-run relationship, thus providing unbiased estimates (Pesaran et. al, 2001). So, it is possible to express a simplified ARDL model as;

$$\Delta y_{t} = \beta_{1} + \delta_{1} y_{t-1} + \delta_{2} x_{t-1} + \delta_{3} z_{t-1} + \eta_{1} \sum_{i=1}^{n} \Delta y_{t-1} + \eta_{2} \sum_{i=1}^{n} \Delta x_{t-1} + \eta_{3} \sum_{i=1}^{n} \Delta z_{t-1} + \mu_{t}$$
(1)

Where, the long-run coefficients are $(\delta_1, \delta_2, \delta_3)$ whose sum of which corresponds to the ECM and, η_1, η_2 ,

 η_3 are the short-run coefficients. The generalized ARDL model for investigating the effect on growth of the stock market is specified as follows;

$$\Delta \ln(Y)_{t} = \alpha_{0} + \beta_{1} \Delta \ln(MCAP)_{t-1} + \beta_{2} \Delta \ln(GFCF)_{t-1} + \beta_{3} \Delta \ln(M3)_{t-1} + \beta_{4} \Delta \ln(Y)_{t-1} + \beta_{5} \Delta \ln(TURNOVER)_{t-1} + \beta_{6} \Delta \ln(VALTRD)_{t-1} + \phi_{7} \log(MCAP)_{t} + \phi_{8} \log(GFCF)_{t}$$
(2)
+ $\phi_{9} \log(M3)_{t} + \phi_{10} \log(TURNOVER)_{t} + \phi_{11} \log(VALTRD)_{t} + \mu_{t}$

where μ_t the stochastic term, t-1 is the lagged period, Δ designates differencing of variables and $\phi_7 - \phi_{11}$

are long run coefficients.

The unrestricted ECM is formulated as part of the ARDL approach to define the long run co-integration, taking into account each variables to estimate the best-fit model which is shown in the matrix below;

$$\begin{bmatrix} \Delta \ln(Y)_{t} & \\ \Delta \ln(MCAP)_{t} & \\ \Delta \ln(GFCF)_{t} & \\ \Delta \ln(M3)_{t} & \\ \Delta \ln(TURNOVER)_{t} \\ \Delta \ln(VALTRD)_{t} \end{bmatrix} = \begin{bmatrix} \varphi_{1} \\ \varphi_{2} \\ \varphi_{3} \\ \varphi_{4} \\ \varphi_{5} \\ \varphi_{6} \end{bmatrix} + \begin{bmatrix} \ln(Y)_{t-1} & \\ \ln(MCAP)_{t-1} \\ \ln(GFCF)_{t-1} \\ \ln(M3)_{t-1} \\ \ln(TURNOVER)_{t-1} \end{bmatrix} \begin{bmatrix} \sigma_{11} & \sigma_{12} & \sigma_{13} & \sigma_{14} & \sigma_{15} & \sigma_{16} \\ \sigma_{21} & \sigma_{22} & \sigma_{23} & \sigma_{24} & \sigma_{25} & \sigma_{26} \\ \sigma_{31} & \sigma_{32} & \sigma_{33} & \sigma_{34} & \sigma_{35} & \sigma_{36} \\ \sigma_{41} & \sigma_{42} & \sigma_{43} & \sigma_{44} & \sigma_{45} & \sigma_{46} \\ \sigma_{51} & \sigma_{52} & \sigma_{53} & \sigma_{54} & \sigma_{55} & \sigma_{56} \\ \sigma_{61} & \sigma_{62} & \sigma_{63} & \sigma_{64} & \sigma_{65} & \sigma_{66} \end{bmatrix}$$

$$+ \sum_{h=1}^{q} \begin{bmatrix} \upsilon_{11h} & \upsilon_{12h} & \upsilon_{13h} & \upsilon_{14h} & \upsilon_{15h} & \upsilon_{16h} \\ \upsilon_{21h} & \upsilon_{22h} & \upsilon_{23h} & \upsilon_{24h} & \upsilon_{25h} & \upsilon_{26h} \\ \upsilon_{31h} & \upsilon_{32h} & \upsilon_{33h} & \upsilon_{34h} & \upsilon_{35h} & \upsilon_{36h} \\ \upsilon_{31h} & \upsilon_{32h} & \upsilon_{33h} & \upsilon_{34h} & \upsilon_{35h} & \upsilon_{36h} \\ \upsilon_{51h} & \upsilon_{52h} & \upsilon_{53h} & \upsilon_{54h} & \upsilon_{55h} & \upsilon_{56h} \\ \upsilon_{61h} & \upsilon_{62h} & \upsilon_{63h} & \upsilon_{64h} & \upsilon_{65h} & \upsilon_{66h} \end{bmatrix} \begin{bmatrix} \Delta \ln(Y)_{t-h} \\ \Delta \ln(MCAP)_{t-h} \\ \Delta \ln(VALTRD)_{t-h} \end{bmatrix} + \begin{bmatrix} \mu_{1t} \\ \mu_{2t} \\ \mu_{3t} \\ \mu_{4t} \\ \mu_{5t} \\ \mu_{6t} \end{bmatrix}$$

where $\varphi_1 - \varphi_6$ represents constant terms, $\sigma_{11} - \sigma_{66}$ represents long run coefficients, $\upsilon_{11} - \upsilon_{66}$ represents short run coefficients, and Δ for the first difference. The bound test is conducted using the F-tests to evaluate the long-run relationship between the variables. In equation 3, the null hypothesis of no-cointegration is;

$$H_{0} = \begin{bmatrix} \sigma_{11} = & \sigma_{12} = & \sigma_{13} = & \sigma_{14} = & \sigma_{15} = & \sigma_{16} \\ \sigma_{21} = & \sigma_{22} = & \sigma_{23} = & \sigma_{24} = & \sigma_{25} = & \sigma_{26} \\ \sigma_{31} = & \sigma_{32} = & \sigma_{33} = & \sigma_{34} = & \sigma_{35} = & \sigma_{36} \\ \sigma_{41} = & \sigma_{42} = & \sigma_{43} = & \sigma_{44} = & \sigma_{45} = & \sigma_{46} \\ \sigma_{51} = & \sigma_{52} = & \sigma_{53} = & \sigma_{54} = & \sigma_{55} = & \sigma_{56} \\ \sigma_{61} = & \sigma_{62} = & \sigma_{63} = & \sigma_{64} = & \sigma_{65} = & \sigma_{66} \end{bmatrix} = 0$$

As against the alternative hypothesis is;

$$H_{1} = \begin{bmatrix} \sigma_{11} \neq & \sigma_{12} \neq & \sigma_{13} \neq & \sigma_{14} \neq & \sigma_{15} \neq & \sigma_{16} \\ \sigma_{21} \neq & \sigma_{22} \neq & \sigma_{23} \neq & \sigma_{24} \neq & \sigma_{25} \neq & \sigma_{26} \\ \sigma_{31} \neq & \sigma_{32} \neq & \sigma_{33} \neq & \sigma_{34} \neq & \sigma_{35} \neq & \sigma_{36} \\ \sigma_{41} \neq & \sigma_{42} \neq & \sigma_{43} \neq & \sigma_{44} \neq & \sigma_{45} \neq & \sigma_{46} \\ \sigma_{51} \neq & \sigma_{52} \neq & \sigma_{53} \neq & \sigma_{54} \neq & \sigma_{55} \neq & \sigma_{56} \\ \sigma_{61} \neq & \sigma_{62} \neq & \sigma_{63} \neq & \sigma_{64} \neq & \sigma_{65} \neq & \sigma_{66} \end{bmatrix} = 0$$

According toPesaran et. al (2001), the criteria for the decision making are as follows;

i. The presence of cointegration is verified if F_h > upper bound critical value

ii. If F_h < upper-bound critical value, it confirms that the variables are not cointegrated

iii. If $F_h \leq l \geq$ upper/lower bound critical value, then the cointegration decision is inconclusive

The long run equilibrium is thus modelled as follows;

$$\ln Y_{t} = \omega_{0} + \sum_{k=1}^{m} \alpha_{k} \ln(Y)_{t-k} + \sum_{k=0}^{n} \beta_{k} \ln(MCAP)_{t-k} + \sum_{k=0}^{p} \nu_{k} \ln(GFCF)_{t-k} + \sum_{k=0}^{q} \pi_{k} \ln(M3)_{t-k} + \sum_{k=0}^{r} \chi_{k} \ln(TURNOVER)_{t-k} + \sum_{k=0}^{r} \chi_{k} \ln(VALTRD)_{t-k} + \varepsilon_{t}$$
(4)

The short-run elasticities can be generated by formulating the following error correction;

$$\Delta(\ln Y)_{t} = \omega_{0} + \sum_{k=1}^{m} \alpha_{k} \Delta \ln(Y)_{t-k} + \sum_{k=0}^{n} \beta_{k} \Delta \ln(MCAP)_{t-k} + \sum_{k=0}^{p} \nu_{k} \Delta \ln(GFCF)_{t-k} + \sum_{k=0}^{q} \pi_{k} \Delta \ln(M3)_{t-k} + \sum_{k=0}^{r} \chi_{k} \ln(TURNOVER)_{t-k} + \sum_{k=0}^{r} \chi_{k} \ln(VALTRD)_{t-k} + \varsigma ECT_{t-1} + \psi_{t}$$
(5)

Table 4. Description and nomenclature of variables

| Variable | Variable Definition | |
|--------------------------------------|--|----------|
| Dependent variables | | |
| Economic growth (EGWTH) | An inflation-adjusted indicator reflecting the value of all goods and services produced following Ishioro (2013); Hasan & Barau (2015) by an economy measured in constant price. | |
| Independent variable | | |
| Gross fixed capital formation (GFCF) | Measures the net increase in fixed capital as a percentage of GDP | Positive |
| Money supply (M3) | A measure of money supply which includes M2 as well as large time deposits, retail money market funds, short-term repurchase agreements and high liquid | Positive |

| | assets, emphasizing money more as a store-of-value than as a means of | |
|-----------------------|---|----------|
| | exchange, calculated as a percentage of GDP | |
| Market capitalization | According to Nyasha & Odhiambo (2015, 2016), the aggregate value of the | Positive |
| (MCAP) | firm based on the current share price and the total number of outstanding | |
| | stocks that serves as an instrument enabling the investor assess the returns and | |
| | risk in the share, measured as percent ageof GDP. | |
| Turnover ratio | This ratio represents the proportion of stocks that changes in a fiscal year. This is used to evaluate the company's efficiency on how it uses its assets to | Positive |
| | generate revenue, measured in percentage ponts, following Nyasha & | |
| | Odhiambo (2015, 2016) | |
| Value of stock traded | According to Nyasha & Odhiambo (2015, 2016), this refers to the total | Positive |
| (VALTRD) | number of shares-traded, both domestic and foreign, multiplied by their | |
| | respective matching prices | |
| | | |

Source: Compiled by Author

4. Analysis of Data and Discussion of Results

This section began with the descriptive statistics and correlation properties of variables used. On the premise of the data obtained for Nigeria, the mean for economic growth for the period is approximately 17.31 percent, with maximum and minimum value (18.061) and (16.648) percent respectively. The mean values for market capitalization, turn-over ratio and value of shares traded are (13.122 percent), (1.288 percent) and (9.817 percent), with their corresponding maximum and minimum values of (16.427 and 7.916 percent), (3,526 and -1.203 percent) and (14.680 and 2.631) percent respectively. The *p*-values of the Jarque-Bera indicate that all of the variables are normally distributed as seen in Table 5, Panel B.

| | EGWTH | GFCF | <i>M3</i> | MCAP | TURNOVER | VALTRD |
|-------------|----------|------------------|----------------------|-----------------|----------|----------|
| | | Panel- A: De | scriptive statistic | s (level value) | | |
| Mean | 27126109 | 31.362 | 5757666 | 3770831 | 6.891 | 437477.8 |
| Maximum | 28701907 | 54.948 | 25079721 | 13619906 | 34.000 | 2375619 |
| Minimum | 16997518 | 14.168 | 22299.24 | 2743.100 | 0.300 | 13.900 |
| Std. Dev. | 18640860 | 13.476 | 7753023 | 4691060 | 6.832 | 626633.6 |
| Jarque-Bera | 4.099 | 2.751 | 7.862 | 5.077 | 62.212 | 18.477 |
| Probability | 0.128 | 0.252 | 0.019 | 0.078 | 0.000 | 0.000 |
| | ŀ | Panel- B: Descri | iptive statistics (r | atural log valu | e) | |
| Mean | 17.310 | 3.346 | 13.892 | 13.122 | 1.288 | 9.817 |
| Maximum | 18.061 | 4.006 | 17.037 | 16.427 | 3.526 | 14.680 |
| Minimum | 16.648 | 2.651 | 10.012 | 7.916 | -1.203 | 2.631 |
| Std. Dev. | 0.491 | 0.464 | 2.333 | 2.870 | 1.345 | 4.155 |
| Jarque-Bera | 3.386 | 3.058 | 2.648 | 3.155 | 3.415 | 3.720 |
| Probability | 0.183 | 0.216 | 0.266 | 0.206 | 0.181 | 0.155 |
| | | Panel | C: Correlation | Matrix | | |
| EGWTH | 1 | | | | | |
| GFCF | 0.911** | 1 | | | | |
| М3 | 0.959** | 0.797** | 1 | | | |
| MCAP | 0.950** | 0.849** | 0.924** | 1 | | |
| TURNOVER | 0.519** | 0.679** | 0.366** | 0.534** | 1 | |
| VALTRD | 0.725** | 0.715** | 0.660** | 0.828** | 0.838** | 1 |

Table 5. Summary statistics for period (1985-2018)

Source: Author's Computation using E-views 10

N.B: EGWTH = Economic Growth; GFCF = Gross Fixed Capital Formation; M3: Broad Money Supply; MCAP = Market Capitalization; TURNOVER = Turnover ratio; VALTRD = Value of shares traded; Std. Dev. = Standard Deviation; ** signifies significant at 5% level

| | At le | evel | I st Dif | | |
|-----------|------------------|------------------|---------------------|-------------------|----------|
| | ADF (prob.) | PP (prob.) | ADF (prob.) | PP (prob.) | Decision |
| EGWTH | -3.861 (0.025)** | -3.796 (0.029)** | - | - | I(0) |
| GFCF | -3.194 (0.102) | -0.762 (0.816) | -6.687 (0.000)** | -10.009 (0.000)** | I(1) |
| <i>M3</i> | 1.383 (0.999) | 1.388 (0.999) | -4.946 (0.001)** | -4.960 (0.001)** | I(1) |
| MCAP | -2.778 (0.214) | -2.778 (0.214) | -6.465 (0.000)** | -14.519 (0.000)** | I(1) |
| TURNOVER | -2.627 (0.271) | -2.627 (0.271) | -5.511 (0.000)** | -7.631 (0.000)** | I(1) |
| VALTRD | -2.038 (0.269) | -1.912 (0.322 | -5.954 (0.000)** | -10.252 (0.000)** | I(1) |

Table 6. Stationarity Test

Source: Author's Computation using E-views 10

N.B: ** signifies significant at 5% level

The stationarity test in Table 6 indicates mixed order of integration. Economic growth is stationary at level, I(0), whereas at first difference, other variables (capital formation, money supply, market capitalization, turnover ratio and value traded) are integrated at order one, I(1). A mixed order of integration allows the adoption of the ARDL bounds test, based on the aforementioned result, to capture the long-run cointegration among the variables, Pesaran et. al (2001).

Table 7. ARDL Bound test results [Long run Co-integration]

| Model Specificati | on | F-stat | | Lags | | Co-integration | n |
|--|-----------------|---------|------------|------|------|----------------|------|
| $\Delta EGWTH = \left\{ F_{EGWTH} (EGWTH_T) \mid MC \right\}$ | AP, GFCF, M3) | 6.725** | Stability | 4 | | Accepted | |
| $\Delta EGWTH = \left\{ F_{EGWTH} (EGWTH_T) TURNOPPORTUNAL TURPORTUNAL TURPORTUNAL TURPORTUNA T$ | OVER, GFCF, M3) | 7.565** | Depth | 1 | | Accepted | |
| $\Delta EGWTH = \left\{ F_{EGWTH} (EGWTH_T) VALT \right\}$ | TRD, GFCF, M3) | 7.130** | Efficiency | 4 | | Accepted | |
| | K | 1 | !% | 5% | | 10% | |
| Critical Value | | I(0) | I(1) | I(0) | I(1) | I(0) | I(1) |
| Pesaran et. al (2001) | 3 | 4.29 | 5.61 | 3.23 | 4.35 | 2.72 | 3.77 |

Source: Author's Computation using E-views 10; N.B: ** signifies significant at 5% level; Lag length section criterion = Akaike Information Criterion (AIC)

Three (3) ARDL bound test are executed from table 7, showing their respective computed F-statistics. Results indicate a long run relationship across all three models. This means that the no cointegration hypothesis against its alternative is dismissed at c.v (1%, 5% and 10%). The computed F-statistic of 6.725 is higher than the upper bound critical value of 4.35 at 5%, indicating a long-run relationship between market capitalization, gross fixed capital formation, money supply and growth. Therefore, the finding is replicated for the other two models which signifies the depth and efficiency of the interaction of the stock market with growth as revealed from their computed F-statistic values (7.565 and 7.130) respectively. Based on the results shown in table 7, it is evident that stock market development has a long run effect on growth in Nigeria.

| | Stock-Market Development long run empirical assessment | | | | |
|--------------------------------|--|------------------|-----------------|--|--|
| | Equation 1 | Equation 2 | Equation 3 | | |
| | ARDL (1,0,0,1) | ARDL (1,0,0,1) | ARDL (1,0,0,1) | | |
| $\Delta \ln EGWTH_{t-1}$ | 1.168 (4.249)** | 0.917 (11.905)** | 0.985 (3.523)** | | |
| $\Delta \ln MCAP_{t-1}$ | 0.022 (0.797) | - | - | | |
| $\Delta \ln TURNOVER_{t-1}$ | - | 0.020 (1.628) | - | | |
| $\Delta \ln VALTRD_{t-1}$ | - | - | 0.006 (0.496) | | |
| $\Delta \ln \text{GFCF}_{t-1}$ | 0.017 (0.143) | -0.052 (-0.693) | -0.010 (-0.097) | | |
| $\Delta \ln M3_{t-1}$ | 0.061 (0.541) | -0.001 (-0.086) | 0.082 (0.637) | | |
| С | 2.630 (0.638) | 1.634 (1.138) | 4.297 (1.037) | | |

Table 8. Coefficient of long-run ARDL Model (period 1985-2018)

Source: Author's Computation using E-views 10; N.B: ** signifies significant at 5% level

Empirical results show that growth was positively influenced by the first time lag of the three proxies, albeit statistically insignificant at 5% level. In addition, the long run result reveals that the stability, depth and efficiency effect on growth is lacking. These findings are consistent with empirical studies of (Karim et. al 2017; Okoro G.E; 2016, Sinha et. al; 2015, Haque; 2013; Afolabi, 2015 and Echekoba et. al, 2013). This signifies the need for a proper functioning of the Nigerian stock market to fuel economic growth through capital accumulation and accelerated industrialization. The stock market plays a key role in market-based financial development with asset liquidity, long-run capital adequacy for investment, and efficient use of resources. As a result, a well-functioning stock market remains a crucial indicator of macroeconomic development, attracting long term investment from domestic and foreign investors and thereby playing a key role in accelerating industrialization (Pohoata et. al, 2016; Coskun et. al, 2017.

As for the money supply, empirical findings show that it was positively related to growth in equations 1 and 3, although statistically insignificant to growth as seen in table 8. This illustrates the fact that the financial sector has little or no innovation that can stimulate money supply while at the same time minimize investment risk in the economy. The findings are reinforced by other empirical studies (Qamruzzanian and Wei, 2017; Bara et. al, 2016; Bara & Mudzingini, 2016), while gross fixed capital formation shows a positive relationship to growth in equation one only, howbeit, statistically insignificant.

The short run dynamic (ECM_ARDL) model is thus analyzed and described in table 9 based on the confirmed long run relationship.

| | Equation 1 | Equation 2 | Equation 3 |
|--|--------------------|-------------------|-------------------|
| | ARDL (1,0,0,1) | ARDL (1,0,0,1) | ARDL (1,0,0,1) |
| $\Delta lnEGWTH_{t-1}$ | 1.364 (2.553)** | 1.165 (3.643)** | 1.463 (0.015)** |
| $\Delta \ln MCAP_{t-1}$ | 0.019 (0.717) | - | - |
| $\Delta \operatorname{InTURNOVER}_{t-1}$ | - | -0.003 (-0.209) | - |
| $\Delta \ln VALTRD_{t-1}$ | - | - | 0.011 (0.986) |
| $\Delta \ln GFCF_{t-1}$ | 0.051 (0.448) | 0.018 (0.273) | 0.075 (0.831) |
| $\Delta \ln M3_{t-1}$ | 0.077 (0.789) | 0.032 (0.599) | 0.065 (0.786) |
| ECT _{t-1} | -1.121 (-1.899)*** | -1.001 (-2.625)** | -1.575 (-3.628)** |
| | Diag | nostic tests | |
| \mathbf{R}^2 | 0.717 | 0.653 | 0.816 |
| D.W | 1.875 | 2.209 | 2.289 |
| Heteroscedasticity | 0.390 (0.537) | 0.313 (0.580) | 1.218 (0.279) |
| Normality Test | 0.219 (0.896) | - | 1.389 (0.499) |
| Reset Test | 0.110 (0.746) | 0.020 (0.886) | 0.091 (0.768) |

Table 9. Short-run Dynamic (ECM-ARDL) result

Source: Author's Computation using E-views 10; N.B: ** signifies significant at 5% level

The respective speed-of-adjustment to the long run equilibrium is shown in table 9. The result shows that the ECT for each given equation is negative (-1.121, -1.001 and -1.575) and statistically significant. This means that any previous period shocks are to be adjusted at speeds of (112%, 100% and 157%) respectively in the long run. Furthermore, all indicators except the turnover ratio showed positive effects on growth, howbeit statistically insignificant. This suggests that capital adequacy has not produced the desired economic impact in delivering a significant long run effect for the Nigerian economy. For all three equations, financial innovation (M3) has a positive effect on growth, albeit statistically insignificant. This means that, to some degree, money supply in the economy has supported in recuperating output levels and reduce cost, but has not been successfully in the short-run, but can have long run economic effects. Capital formation showed a positive but marginal effect on growth, indicating theoretically that money supply in the economy, either in the form of capital formation or investment, would stimulate growth by increasing economic activity. Diagnostic tests were performed to assess the validity of the model, including the heterscedasticity test, the test of normality and Ramsey-Reset test (Pagan and Hall (1983). It shows that all statistics are significantly higher than 5 percent, indicating that there is no heteroscedasticity problem and first order serial correlation, usually a desirable indicator of a stable econometric model.

5. Conclusion and Recommendations

This study critically investigated the link between stock market development and economic growth in Nigeria for the period, 1985-2018. Reviewed current and existing literature revealed quite a few empirical studies on the subject matters in developed, emerging and developing economies. There have been a number of studies, but a few have been performed on the relationship between stock market development and economic growth has been conducted (Osamwonyi & Kasimu (2013); Edame et. al (2013); Echekoba et. al. (2013); Okoye & Nwisienyi (2013) and Ishioro, 2013) among others and a limited number of others focused on stock market-led growth. Seeing the void in existing studies, the study explored a new dimension in stock market development along with financial innovation, where money supply (M3) was used to capture innovation-effect on growth. The study adopted the ARDL Bound test methodology, in order to encapsulate the long run relationship between stock market development and economic growth.

The bounds test demonstrates the existence of a long run relationship for all three (3) models tested. The findings clearly support a long-run relationship between stock-market development and Nigeria's economic growth. In explaining the long- and short run elasticities, findings revealed that stock market development metrics have positive long- and short run impact on economic growth, although insignificantly. This implies that the development of Nigeria's stock market will potentially improve short- and long run growth of the economy by increased depth, stability and efficient financial institutions, capital accumulation and long term capital adequacy. Financial innovation, proxied by M3, had positive impact, albeit statistically insignificant, on economic growth both in the long and short run periods. This shows that stock market financial innovation is expanding financial services by creating new and accelerated institutions, assets and services that eventually lead to accelerated growth.

As far as financial innovation and its effect on growth, innovation is concerned; innovation plays a crucial role in the development of the stock market, as it creates room for the expansion financial activities in the economy by developing new mechanisms and forms of financial institutions. Therefore, policies aimed at making Nigeria's capital market more financially innovation-oriented should therefore be implemented in order to provide large number of households and firms with integrated services, which then can then contribute to the growth cycle.

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